1.5 Variable Transformations

Let the RV X have pdf f(x) and let h(x) denote some function of X. (e.g. if x is the colour of a star, then h(x) could be the temperature). H = h(x) is itself a RV with pdf g(h), say. How are f(x) and g(h) related? We have that

$$f(x) dx = \operatorname{prob}(x < X < x + dx)$$

and we require to find g(h) such that

$$g(h) dh = \operatorname{prob}(h < H < h + dh)$$

Suppose first that h(x) is one-to-one, i.e. x maps to a unique h and vice versa. Hence the inverse function x = x(h) exists, and we can write f as a function of h, i.e.

$$f(x) \equiv f(x(h))$$

This function is not g(h), however, since we also have to transform the infinitesimal dx (just as with changing variables in integration). Thus

$$dx = \left| \frac{dx}{dh} \right| dh$$

where the modulus is required because probability is never negative. Combining the above expressions:-

$$f(x) dx = f(x(h)) \left| \frac{dx}{dh} \right| dh = g(h) dh$$

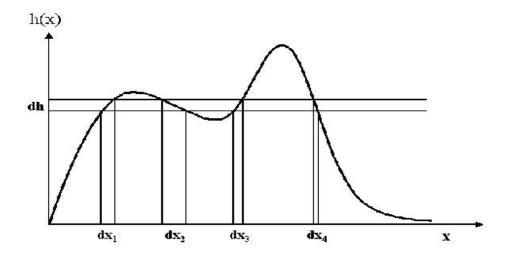
If h(x) is not one-to-one then we must sum over all values of x for which h(x) = h, or more precisely over the small intervals, dx_i , corresponding to dh (see Figure 7). Thus

$$\operatorname{prob}(h < H < h + dh) = \operatorname{prob}(x_1 < x < x_1 + dx_1) + \operatorname{prob}(x_2 < x < x_2 + dx_2) + \operatorname{prob}(x_3 < x < x_3 + dx_3) + \dots$$

It then follows that

$$g(h) dh = \sum_{h(x_i)=h} f(x_i(h)) \left| \frac{dx}{dh} \right|_{x_i(h)} dh$$

Figure 7: Variable transformation when h is not one-to-one



1.6: Probability Integral Transform

One variable transformation merits special consideration. Suppose X has PDF f(x) and CDF F(x). Define $h(x) \equiv F(x)$, which is one-to-one. Then:-

$$g(h) dh = f(x(h)) \left| \frac{dx}{dh} \right| dh$$

$$= f(x(h)) \left| \frac{dh}{dx} \right|^{-1} dh$$
(1)

Since h(x) = F(x), dh/dx = f(x). Thus

$$g(h) dh = \frac{f(x)}{f(x)} dh = 1.dh$$

i.e. the pdf of H is the uniform distribution, U(0,1), (since 0 < F(x) < 1).

This important result shows that we can always transform the PDF of any RV into the simple form of U(0,1), **provided** we know the CDF of the original RV. This approach can be used in generating random numbers numerically (see e.g. Numerical Recipes, Chap 17.).

1.7: Multivariate Distributions

Thus far we have considered only the properties of distributions of a single (univariate) RV. We now extend to the **multivariate** case of two or more RVs.

1.7.1 : Joint PDF

The **joint PDF** of two RVs, X_1 and X_2 is $p(x_1, x_2)$. Then,

$$\operatorname{Prob}(a_1 < X_1 < b_1 \text{ and } a_2 < X_2 < b_2) = \int_{a_1}^{b_1} \int_{a_2}^{b_2} p(x_1, x_2) \, dx_1 dx_2$$

Extension to more than two RVs is carried out in the obvious way.

1.7.2: Marginal Distributions

The **marginal PDF**, $p_1(x_1)$ of X_1 is defined by

$$p_1(x_1) = \int_{-\infty}^{\infty} p(x_1, x_2) dx_2$$

and is a PDF in the usual sense that

- 1. $p_1(x_1) > 0$, for all x_1
- 2. Prob $(a < X_1 < b) = \int_a^b p_1(x_1) dx_1$
- 3. $\int_{-\infty}^{\infty} p_1(x_1) dx_1 = 1$

Similarly, the marginal PDF of X_2 is

$$p_2(x_2) = \int_{-\infty}^{\infty} p(x_1, x_2) dx_1$$

In general, given any multivariate PDF, we may find the marginal PDF of any subset of the $X_1, ..., X_n$ by integrating over all other variables. e.g.

$$p_{13}(x_1, x_3) = \int_{-\infty}^{\infty} ... \int_{-\infty}^{\infty} p(x_1, ..., x_n) dx_2 dx_4 dx_5 ... dx_n$$

1.7.3 : Conditional Distributions

Consider the joint PDF, $p(x_1, x_2)$, of X_1 and X_2 . Suppose we observe X_1 to have the value x_1 , but do not observe X_2 . We want a function that describes the PDF of X_2 , given that $X_1 = x_1$ (usually simply stated as 'given x_1 '). This function is known as the **conditional** PDF of X_2 , written as $p(x_2|x_1)$, and defined by

$$p(x_2|x_1) = \frac{p(x_1, x_2)}{p_1(x_1)}$$

i.e. the conditional PDF is obtained by dividing the joint PDF of X_1 and X_2 by the marginal PDF of x_1 (provided $p_1(x_1) \neq 0$). Similarly

$$p(x_1|x_2) = \frac{p(x_1, x_2)}{p_2(x_2)}$$

Note that we can write

$$p(x_1, x_2)$$
 = $p(x_1|x_2)p_2(x_2)$
 = $p(x_2|x_1)p_1(x_1)$

This is known as **Bayes' formula**.

Extension to more than 2 RVs is again straightforward. For example,

$$p(x_1, x_3 | x_2, x_4) = \frac{p(x_1, x_2, x_3, x_4)}{p_{24}(x_2, x_4)}$$

1.8 : Statistical Independence

If the conditional PDF of X_2 given x_1 does not depend on x_1 , this means that X_1 and X_2 are statistically independent, since the observed value of X_2 is unaffected by the observed value of X_1 .

Equivalently, X_1 and X_2 are independent if and only if the joint PDF of X_1 and X_2 can be written as the product of their marginal PDFs, i.e.

$$p(x_1, x_2) = p_1(x_1) p_2(x_2)$$

Again, we extend in the obvious way. The RVs $X_1,...,X_n$ are **mutually independent** if and only if their joint PDF can be written as the product of their marginal pdfs. i.e.

$$p(x_1, x_2, ..., x_n) = p_1(x_1)p_2(x_2)...p_n(x_n)$$